

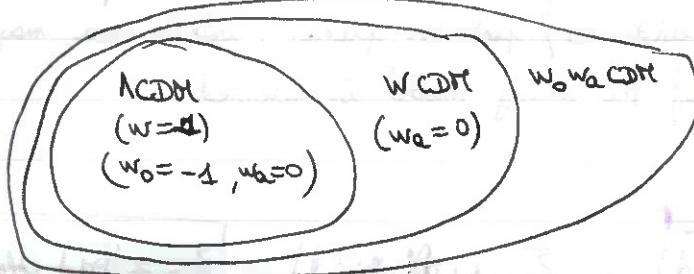
## Bayesian model comparison.

- Basics:  $B_{12} = \frac{\pi(d|\mathcal{M}_1)}{\pi(d|\mathcal{M}_2)}$  assumes implicitly  $\pi(\mathcal{M}_1) = \pi(\mathcal{M}_2)$  ratio of evidences. ( $\neq$  ratio of likelihoods) in frequentist stats.
- Bayes factor.
- $$\pi(d|\mathcal{M}_1) = \int_{\text{evidence}} \underbrace{\pi(d|\theta, \mathcal{M}_1)}_{\text{prior}} \underbrace{\pi(\theta|\mathcal{M}_1)}_{\text{likelihood prior}} d\theta.$$
- = marginal likelihood  $\pi(\mathcal{M}_1)$

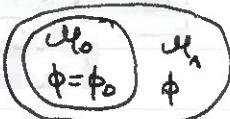
## Nested models and the Savage-Dickey Density Ratio (SSDR):

Example:

nested  
models -



suppose



$$B_{01} = \frac{\pi(d|M_0)}{\pi(d|M_1)}$$

$$\pi(d|M_0) = \int \pi(d|\Psi, M_0) \pi(\Psi|M_0) d\Psi.$$

continuity condition:  $\pi(\Psi|M_0) = \pi(\Psi|\phi=\phi_0, M_1).$

$$\Rightarrow \pi(d|M_0) = \int \pi(d|\Psi, \phi=\phi_0, M_1) \pi(\Psi|\phi=\phi_0, M_1) d\Psi = \pi(d|\phi=\phi_0, M_1)$$

$$= \frac{\pi(\phi=\phi_0|d, M_1) + \pi(d|M_1)}{\pi(\phi=\phi_0|M_1)}$$

$$\Rightarrow B_{01} = \frac{\pi(\phi=\phi_0|d, M_1)}{\pi(\phi=\phi_0|M_1)} \quad (\text{SSDR}). \quad \text{These are usually low-d probabilities.}$$

## Bayesian model selection as a decision analysis

$\{\mathcal{M}_k\}_{1 \leq k \leq N_m}$ : set of models     $\{a_k\} = \{ \text{"selecting } \mathcal{M}_k \} \}_{1 \leq k \leq N_m}$ : set of actions.

$\Rightarrow$  Bayesian decision theory says: take action  $a^*$ :

$$a^* = \underset{k}{\operatorname{argmax}} \quad U(a_k | d)$$

$$U(a_k | d) = \sum_{h=1}^H G(a_k | M_h) \pi(M_h | d)$$

With this formalism, instead of the posterior odds of pairs of models,

$$P_{hk} = \pi(M_h | d) / \pi(M_{h'} | d), \text{ one can quote the ratios}$$

$$U_{hk} = U(M_h | d) / U(M_{h'} | d) = \text{Bayes utility factors.}$$

Simplest choice: 0-1 gain functions:  $G(a_k | M_h) = \delta_{hk}^{a_k} \Rightarrow U(a_k | M_h) = P(M_h | d)$ .

$P_{hk} = U_{hk}$ , we are back to usual Bayesian model comparison.

More generally: one may want to favour models for reasons that are not encoded in their evidences / posterior proba., and/or one may not want to "lose everything" if the wrong model is selected.

### • Model averaging:

$$\pi(\theta | d) = \sum_k \pi(\theta | d, M_k) = \underbrace{\sum_k \pi(\theta | d, M_k)}_{\text{posterior of } \theta \text{ in } M_k} \underbrace{\pi(M_k | d)}_{\text{posterior of } M_k}$$

in particular:

$$E(\theta | d) = \sum_k E(\theta | d, M_k) \pi(M_k | d).$$

$$\text{Var}(\theta | d) = \sum_k [ \text{Var}(\theta | d, M_k) + E(\theta | d, M_k)^2 ] \pi(M_k | d) - E(\theta | d)^2$$

for cases where all the  $\pi(M_k)$  are equal:

$$\pi(\theta | d) \propto \sum_k \underbrace{\pi(\theta | d, M_k)}_{\text{posterior of } \theta \text{ in } M_k} \underbrace{\pi(d | M_k)}_{\text{evidence of } M_k}$$

### • Model selection with insufficient summary statistics:

within a model  $S(d)$  is sufficient for parameter  $\theta$  if and only if:

$$\begin{aligned} \pi(\theta | S(d)) &= \pi(\theta | d, S(d)) \iff \pi(d | S(d), \theta) = \pi(d | S(d)). \\ &= f(\theta | d) \end{aligned}$$

$$B_{12} = \frac{\pi(d | M_1)}{\pi(d | M_2)} = \frac{f(d | S(d), M_1)}{f(d | S(d), M_2)} \frac{\pi(S(d) | M_1)}{\pi(S(d) | M_2)} = \frac{\pi(d | S(d), M_1)}{\pi(d | S(d), M_2)} B_{12}^S$$

true Bayes factor for full data

not bounded, a priori!  
approximate Bayes factor from statistics summary.

$S(d)$  is sufficient for comparing  $M_1$  and  $M_2$  if and only if

$$\pi(d | S(d), M_1) = \pi(d | S(d), M_2) \quad (\Rightarrow B_{12} = B_{12}^S) \quad || \text{ Sufficiency for } M_1, M_2 \text{ or even}$$

Danger in pert. for ABC:  $B_{12}^S$  can be biased, and the for comparing  $M_1$  and  $M_2$ .  
approximation error unrelated to the computational effort (Robert et al. 2011)

(Didelot et al. 2011)